

Chapter 1

Spatio-Temporal Kriging

Workshop *DailyMeteo 2014*,
Spatio-Temporal Geostatistics,
2014-06-24

Introduction

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Kriging

- empirical
- metric
- separable
- product-sum
- sum-metric

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block kriging

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University of Muenster

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From a purely statistical perspective, spatial data is multivariate data with special covariates: the coordinates.

Tobler's first law of Geography states [3]:

Everything is related to everything else, but near things are more related than distant things.

We model the earth, but think in maps: locations are projected from a curved surface in 3D to flat 2D space.

Be aware of geographic coordinates and different projections that maintain angles, certain distances or area.

Imagine the following distances between:

- the Fjord of Oslo (59.85 N 10.75 E) and Uppsala (59.85 N 17.63 E) that are at the same latitude:

Degrees: 6.88

Great Circle: 385 km

Rate: 56 km/degree

- the intersections of the Congo river with the equator (0.00 N 18.21 E) and (0.00 N, 25.53 E):

Coordinate Reference System

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Degrees: 7.32

Great Circle:

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Great Circle: 814 km

Rate: 111 km/degree

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To distinguish different projections, a well prepared data set comes with its coordinate reference system (CRS) as metadata.

These are often encoded as

- EPSG-codes (by the European Petroleum Survey Group)
- proj4string

They define how the reference surface (sphere, ellipsoid) is fixed to the real world (called the datum) and how the projection (surface in 3D to 2D plane) is made.

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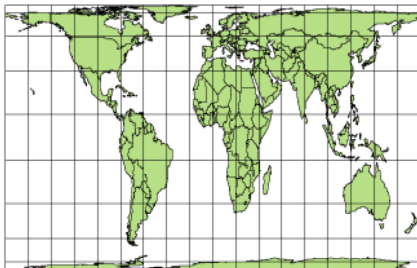
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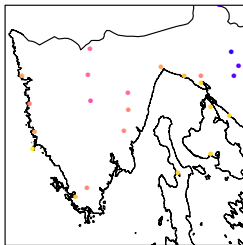
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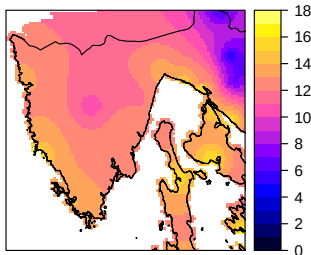


Fields are understood as continuously spreading over space and/or time (e.g. temperature recordings) and typically observed at a set of distinct locations for a series of time steps. Fields are typically illustrated as interpolated maps and modelled as a realisation of a spatial/spatio-temporal random field.

Temp. 2008-09-17



Interpol. Temp. 2008-09-17



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stationarity The process "looks" the same at each location (e.g. mean and variance do not change from east to west)

isotropy The dependence between locations is determined only by their separating distance neglecting the direction (e.g. locations 2 km apart along the north-south axis are as correlated as stations 2 km apart along the east-west axis)

Some tricks exist to weaken these assumptions (e.g. rotating and rescaling coordinates).

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The dependence across space of a random field Z is assessed using a *variogram* γ :

$$\gamma(h) = \frac{1}{2} \mathbb{E} (Z(s) - Z(s+h))^2$$

the empirical estimator looks like

$$\hat{\gamma}(h) = \frac{1}{2|N_h|} \sum_{(i,j) \in N_h} (Z(s_i) - Z(s_j))^2$$

while $N_h = \{(i, j) : h - \epsilon \leq \|s_i - s_j\| \leq h + \epsilon\}$

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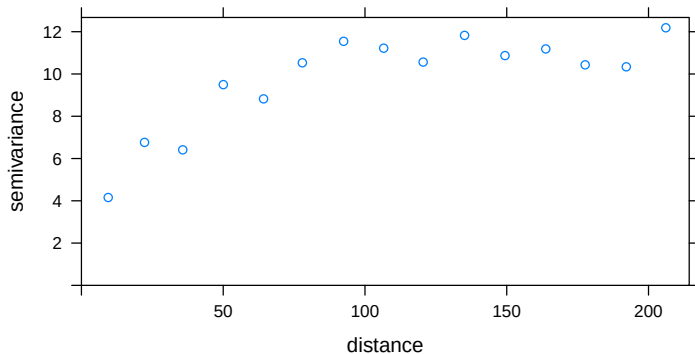
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Variograms II

The *sample variogram* is obtained through

```
vgmTempDay <- variogram(TEMP~1, tempDay)
```

2008-09-17



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And a theoretical *variogram model* can be fitted

```
> head(vgm())
      short                                long
1   Nug                                Nug (nugget)
2   Exp                                Exp (exponential)
3   Sph                                Sph (spherical)
4   Gau                                Gau (gaussian)
5   Exc Exclass (Exponential class)
6   Mat                                Mat (Matern)

> vgmSphHR <- fit.variogram(vgmTempDay,
                           vgm(10,"Sph", 120, 4))

> vgmSphHR
      model    psill    range
1   Nug 3.338608  0.00000
2   Sph 7.737535 98.48023
```

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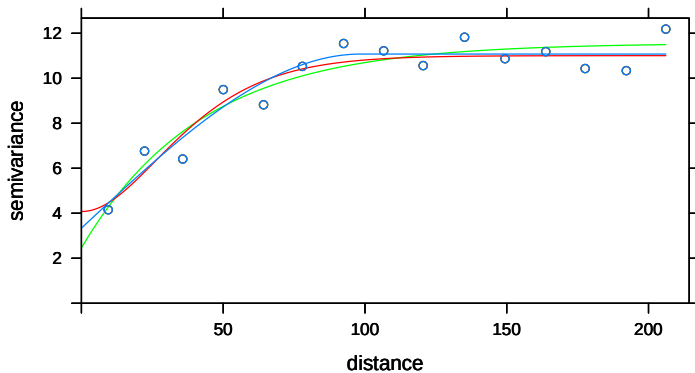
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Certain variogram models can be used to parametrize a covariance matrix for a Gaussian random field over a finite set of locations $s_1, \dots, s_n$:

$$Z \sim \text{Gau}(\boldsymbol{\mu}, \Sigma)$$

while $\Sigma = (\sigma_{ij}^2)_{ij}$ and $\sigma_{ij}^2 = \sigma^2 - \gamma(\|s_i - s_j\|)$, $1 \leq i, j \leq n$ with $\sigma^2 = \text{Var}(Z(s))$, $\boldsymbol{\mu} = (\mu_1, \dots, \mu_n)$.

Predictions can be made using matrix inversion and matrix multiplications.

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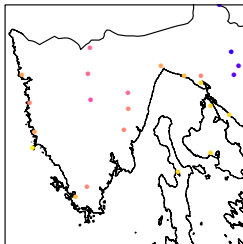
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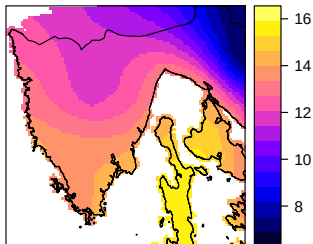
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```
krige(TEMP~1, tempDay, geometry(LST), model=vgmSphHR)
```

Temp. 2008-09-17



Kriged Temp. 2008-09-17



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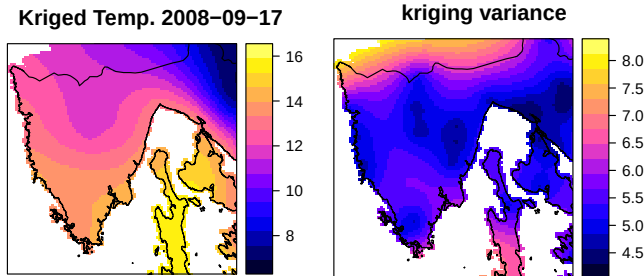
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The model quantifies how *uncertain* it is about the estimates through the kriging variance:



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simple kriging the mean value is known

ordinary kriging prediction based on coordinates

universal kriging prediction based on coordinates and
additional regressors (elevation)

co-kriging the cross-variogram between two variables is as
well exploit (solar radiation)

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Overview of kriging types

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$S \times T$ works as a data structure, but modelling needs to consider special properties of the product of space and time.

direction Today's values influence tomorrow, but will not take effect on yesterday's values.

anisotropy What is the equivalent in terms of dependence of 1 m separation in seconds or minutes?

The easiest way to think of spatio-temporal data is as time slices - but this neglects the temporal dependence.

After modelling temporal trend or periodicities, the residuals might be modelled as a spatio-temporal random field.

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slice wise the easiest adoption is to do interpolation per slice fitting a variogram model for each time slice

pooled the variogram is fitted based on all spatio-temporal data and is used to predict each time slice separately with the same model

evolving models mix the both extremes such that the variogram model adopts to the daily situation (e.g. in terms of overall variability, the sill) but range and the nugget/sill ratio depend on larger data samples.

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Extending the variogram to a twoplace function for spatio-temporal random fields $Z(s, t)$:

$$\gamma(h, u) = E(Z(s, t) - Z(s + h, t + u))^2$$

at any location (s, t) and empirical version

$$\hat{\gamma}(h, u) = \frac{1}{2|N_{h,u}|} \sum_{(i,j) \in N_{h,u}} (Z(s_i, t_i) - Z(s_j, t_j))^2$$

$$\text{while } N_{h,u} = \left\{ (i, j) \left| \begin{array}{l} h - \epsilon_s \leq \|s_i - s_j\| \leq h + \epsilon_s \\ u - \epsilon_t \leq t_i - t_j \leq u + \epsilon_t \end{array} \right. \right\}$$

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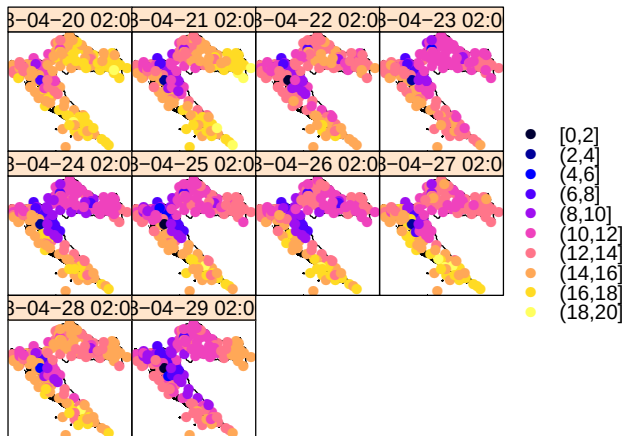
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The idea is the same as in the spatial case: binning of locations according to their separating distance. In the spatio-temporal case, distances are pairs of spatial and temporal distance yielding a variogram surface, not a single line.

```
vgmMeasure <- variogram(TEMP~1, tempST,  
                        tlags=0:7, cutoff=150)  
  
# wireframe:  
plot(vgmMeasure, wireframe=T, scales=list(arrows=F),  
     col.regions=bpy.colors(),zlab=list(rot=90),zlim=c(0,16))  
  
# levelplot:  
plot(vgmMeasure)
```

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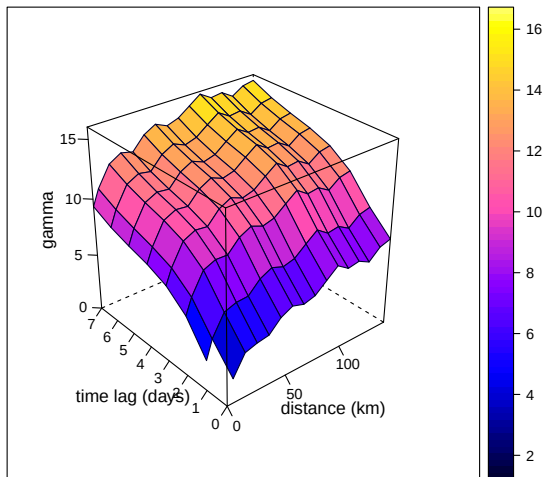
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empirical spatio-temporal variogram surface - wireframe



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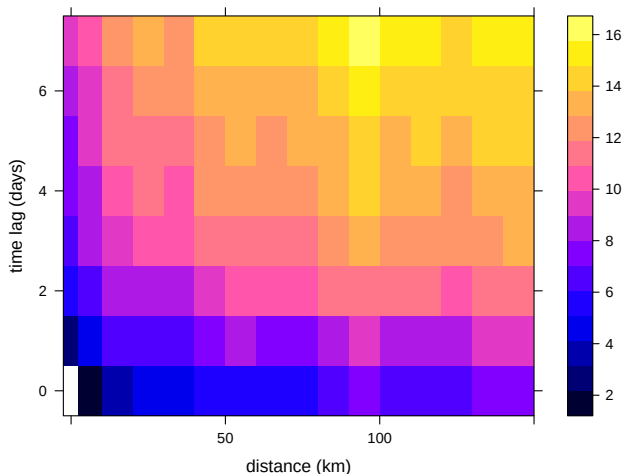
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empirical spatio-temporal variogram surface - levelplot



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The *metric kriging* follows the natural idea of extending the 2-dimensional geographic space into a 3-dimensional spatio-temporal one. In order to achieve an isotropic space, the temporal domain has to be rescaled to match the spatial one (spatio-temporal anisotropy correction κ).

All spatial, temporal and spatio-temporal distances are treated equally resulting in a joint covariance model C_j :

$$C_m(h, u) = C_j(\sqrt{h^2 + (\kappa \cdot u)^2})$$

The variogram evaluates to

$$\gamma_m(h, u) = \gamma_j(\sqrt{h^2 + (\kappa \cdot u)^2})$$

where γ_j is any known variogram including some nugget effect.

```
metricModel <- vgmST("metric",  
                     joint=vgm(12, "Exp", 150, 2),  
                     stAni=50)  
modTempMetric <- fit.StVariogram(vgmMeasure, metricModel)  
attr(modTempMetric, "optim.out")$value # 1.48  
  
plot(vgmMeasure, modTempMetric)  
  
IstriaKrigeMetric <- krigeST(TEMP~1, tempST[,110:120],  
                             targetSTF,  
                             model=modTempMetric)
```

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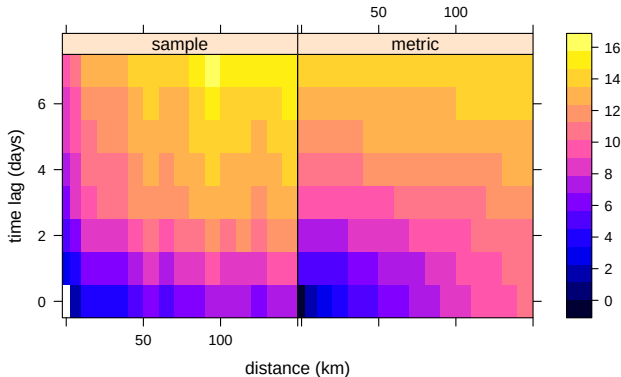
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metric spatio-temporal variogram surface



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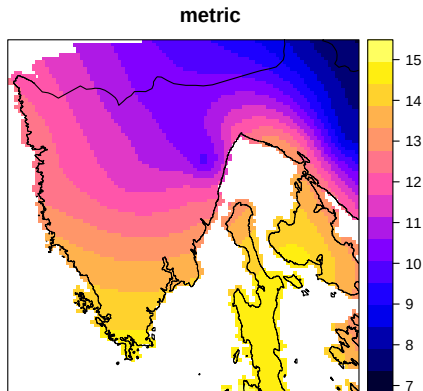
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kriged map for day 115 - metric model



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In space and under the assumptions of isotropy and stationarity, the covariance is a function $C(h)$ of the separating distance h between two locations. A spatio-temporal covariance function is thought of as a function of a spatial and a temporal distance $C(h, t)$.

A *separable covariance function* is assumed to fulfill $C_{sep}(h, u) = C_s(h)C_t(u)$. This is in general a rather strong simplification. Its variogram is given by

$$\gamma_{sep}(h, u) = \text{nug} \cdot \mathbf{1}_{h>0, u>0} + \text{sill} \cdot (\gamma_s(h) + \gamma_t(u) - \gamma_s(h)\gamma_t(u))$$

where γ_s and γ_t are spatial and temporal variograms without nugget effect and a sill of 1. The overall nugget and sill parameters are denoted by "nug" and "sill" respectively.

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```
separableModel <- vgmST("separable",  
                        space=vgm(0.9,"Exp", 147, 0.1),  
                        time =vgm(0.9,"Exp", 3.5, 0.1),  
                        sill=14)  
  
modTempSep <- fit.StVariogram(vgmMeasure, separableModel,  
                             lower=c(10,0,1,0,0))  
attr(modTempSep, "optim.out")$value # Exp-Exp: 0.78  
  
plot(vgmMeasure, modTempSep)  
  
IstriaKrigeSep <- krigeST(TEMP~1, tempST[,110:120],  
                        targetSTF,  
                        model=modTempSep)
```

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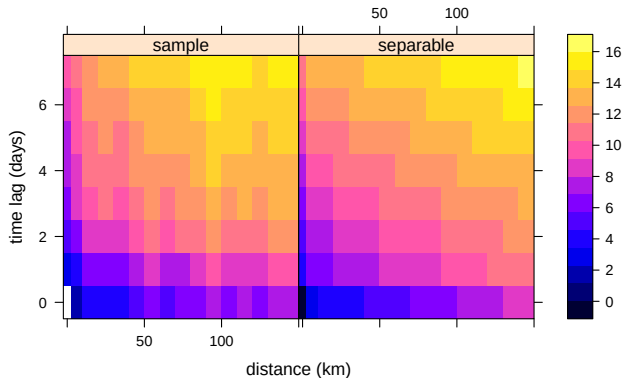
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variogram surface of the product-sum model



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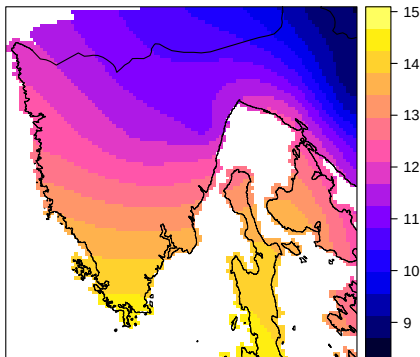
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The *product sum covariance model* extends the simplifying assumption of the separable covariance model to [1]:

$$C_{ps}(h, u) = k_1 C_s(h) + k_2 C_t(u) + k_3 C_s(h) C_t(u)$$

with $k_1 > 0$, $k_2 \geq 0$ and $k_3 \geq 0$ to fulfil the positive-definite condition. The corresponding variogram can be written as

$$\gamma_{ps}(h, u) = \text{nug} \cdot \mathbf{1}_{h>0, u>0} + \gamma_s(h) + \gamma_t(u) - k\gamma_s(h)\gamma_t(u)$$

where γ_s and γ_t are spatial and temporal variograms without nugget effect and in general different sills. The parameter k needs to fulfil $0 < k \leq 1/(\max(\text{sill}_s, \text{sill}_t))$ to let γ_{ps} be a valid model. The overall nugget is denoted by "nug".

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```
prodSumModel <- vgmST("productSum",  
                      space=vgm(10, "Sph", 150, 0),  
                      time= vgm(10, "Exp", 3, 0),  
                      sill=14, nugget=2)  
modTempProdSum <- fit.StVariogram(vgmMeasure, prodSumModel,  
                                  lower=c(0,10,0,1,10,0))  
attr(modTempProdSum, "optim.out")$value # 0.71  
  
plot(vgmMeasure, modTempProdSum)  
  
IstriaKrigeProdSum <- krigeST(TEMP~1, tempST[,110:120],  
                             targetSTF,  
                             model=modTempProdSum)
```

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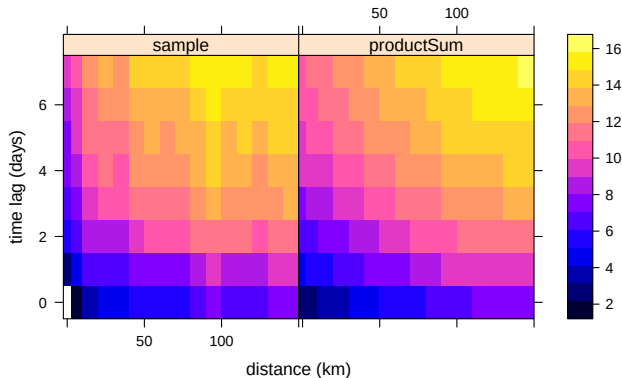
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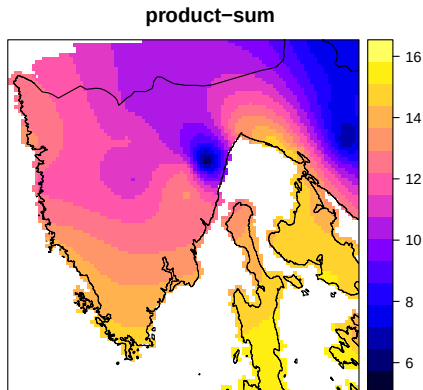
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The *sum-metric covariance model* is given by:

$$C_{sm}(h, u) = C_s(h) + C_t(u) + C_j(\sqrt{h^2 + (\kappa \cdot u)^2})$$

Originally, this model allows for spatial, temporal and joint nugget effects, a simplified version may allow only for a joint nugget. The non-simplified variogram is given by

$$\gamma_{sm}(h, u) = \gamma_s(h) + \gamma_t(u) + \gamma_j(\sqrt{h^2 + (\kappa \cdot u)^2})$$

where γ_s , γ_t and γ_j are spatial, temporal and joint variograms with a separate nugget-effect.

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```
sumMetricModel <- vgmST("sumMetric",  
                        space=vgm( 5, "Exp", 100, 1),  
                        time =vgm( 5, "Mat", 5, 1, kappa=1),  
                        joint=vgm( 5, "Exp", 100, 0),  
                        stAni=20)  
modTempSumMetric <- fit.StVariogram(vgmMeasure, sumMetricModel,  
                                   lower=c(0,10,0,  
                                           0, 1,0,  
                                           0,10,0,  
                                           5))  
attr(modTempSumMetric, "optim.out")$value # 0.44  
  
plot(vgmMeasure, modTempSumMetric)  
  
IstriaKrigeSumMetric <- krigeST(TEMP~1, tempST[,110:120],  
                                targetSTF,  
                                model=modTempSumMetric)
```

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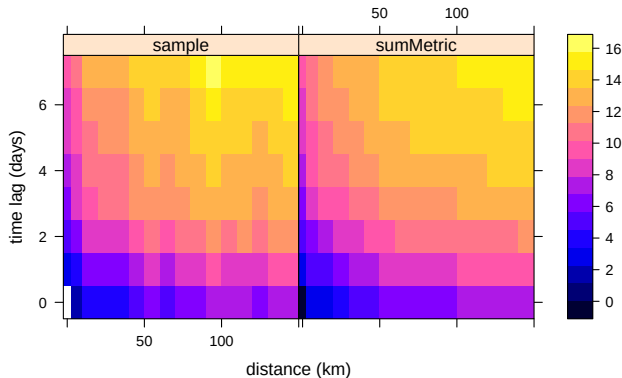
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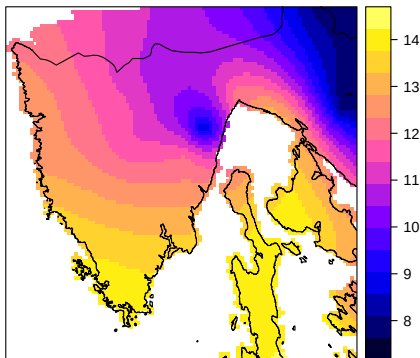
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kriged map for day 115 - sum-metric covariance model

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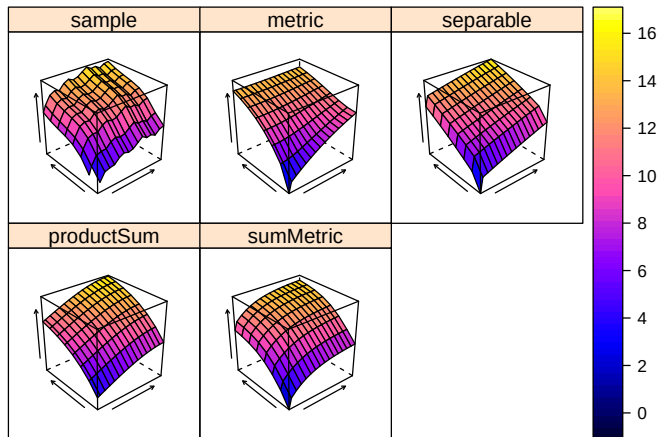
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Purely spatial kriging allows to select the n -nearest neighbours and use only these for prediction.

What does *nearest* mean in a spatio-temporal context?

The idea is to select the most *valuable* locations, i.e. the strongest correlated ones.

Simply set the argument `nmax` and a local neighbourhood of the most correlated values is selected from a larger "metric" neighbourhood.

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```
krigeST(TEMP~1, tempST, nmax=10,  
        targetSTFseries,  
        model=projModTempSumMetric)
```

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In the above scenario and with the presented methods, it is hard to get an uncertainty estimate of the temporally averaged value. Block kriging, with blocks over time, is one way to get such estimates. However, one has to decide on a model beforehand. Here, we will use the metric model again.

Block kriging does not provide estimates for single locations but for areas or volumes. It has the property of providing the correct kriging variance for the block estimate that is typically lower due to the larger area.

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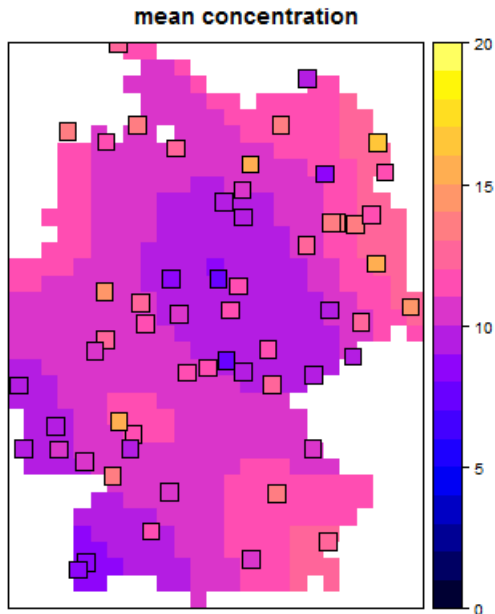
- empirical
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monthly mean concentration - block kriged



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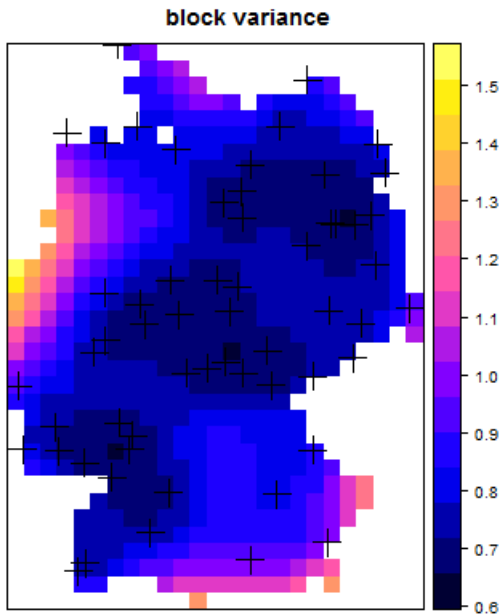
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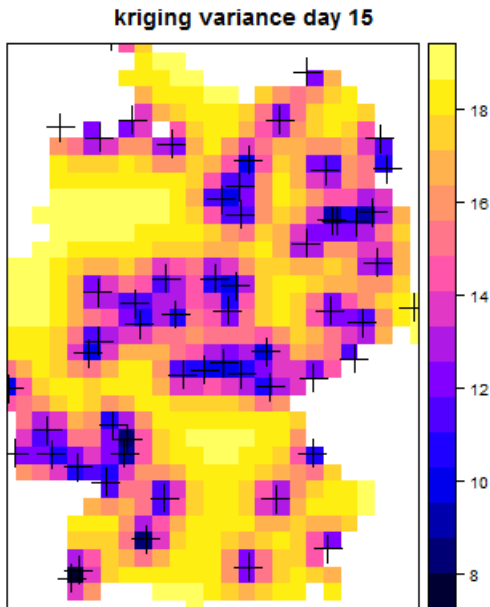
Spatio-Temporal Kriging

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```
tmp_pred <- data.frame(cbind(ger_gridded@coords,15*tmpScale))
colnames(tmp_pred) <- c("x","y","t")
coordinates(tmp_pred) <- ~x+y+t

blockKrige <- krige(PM10~1,
                    air3d[as.vector(!is.na(air3d@data)),],
                    newdata=tmp_pred, model=model3d,
                    block=c(1,1,15*tmpScale))

ger_grid_time@sp@data <- blockKrige@data
```

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References

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- [2] Benedikt Gräler, Lydia E. Gerharz, and Edzer J. Pebesma. Spatio-temporal analysis and interpolation of PM10 measurements in Europe. Technical report, ETC/ACM, 2012.
- [3] W. R. Tobler. A computer movie simulating urban growth in the detroit region. *Economic Geography*, 46:234–240, 1970.

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